# **Selective Randomization Inference** for Adaptive Studies

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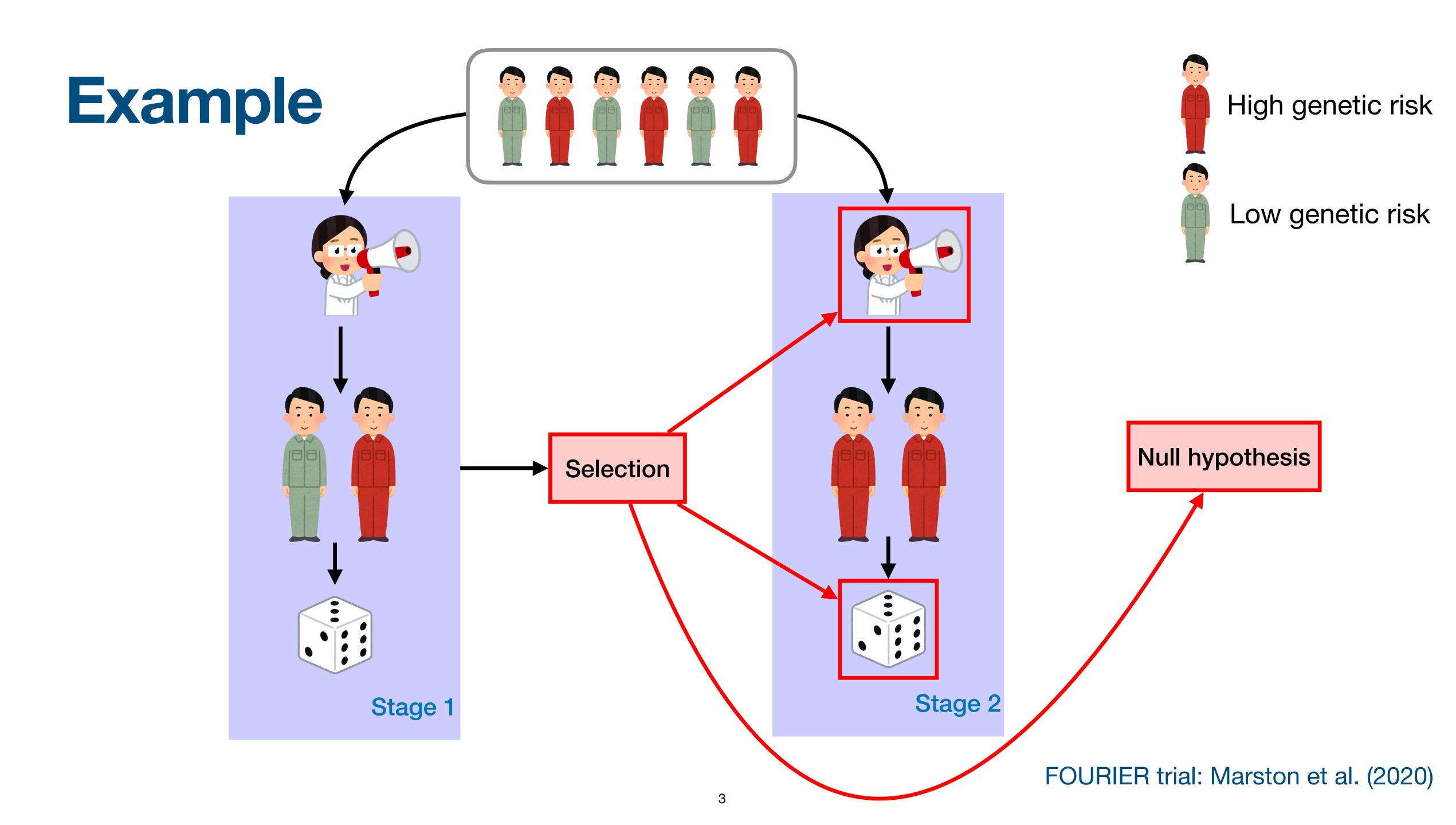
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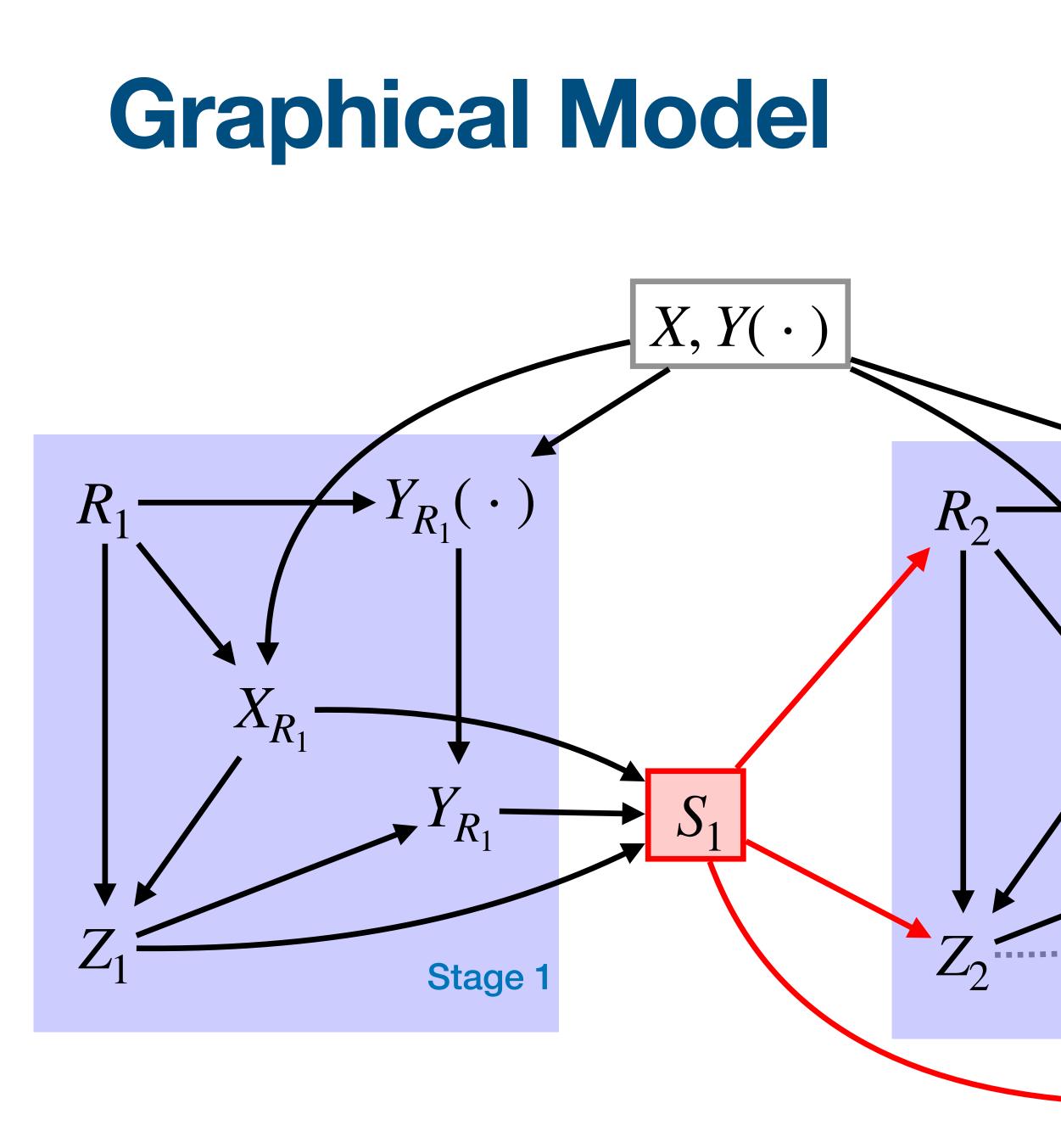
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- Covariates: X
- Potential outcomes:  $Y(\cdot)$
- Recruitment:  $R_k \subseteq [n]$
- Treatments:  $Z_k$

 $S_2$ 

- Observed outcomes: Y = Y(Z)
- Selective choice:  $S_k$
- Short-hand:  $W = (R, X_R, Y_R(\cdot))$

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 $X_{R_2}$ 

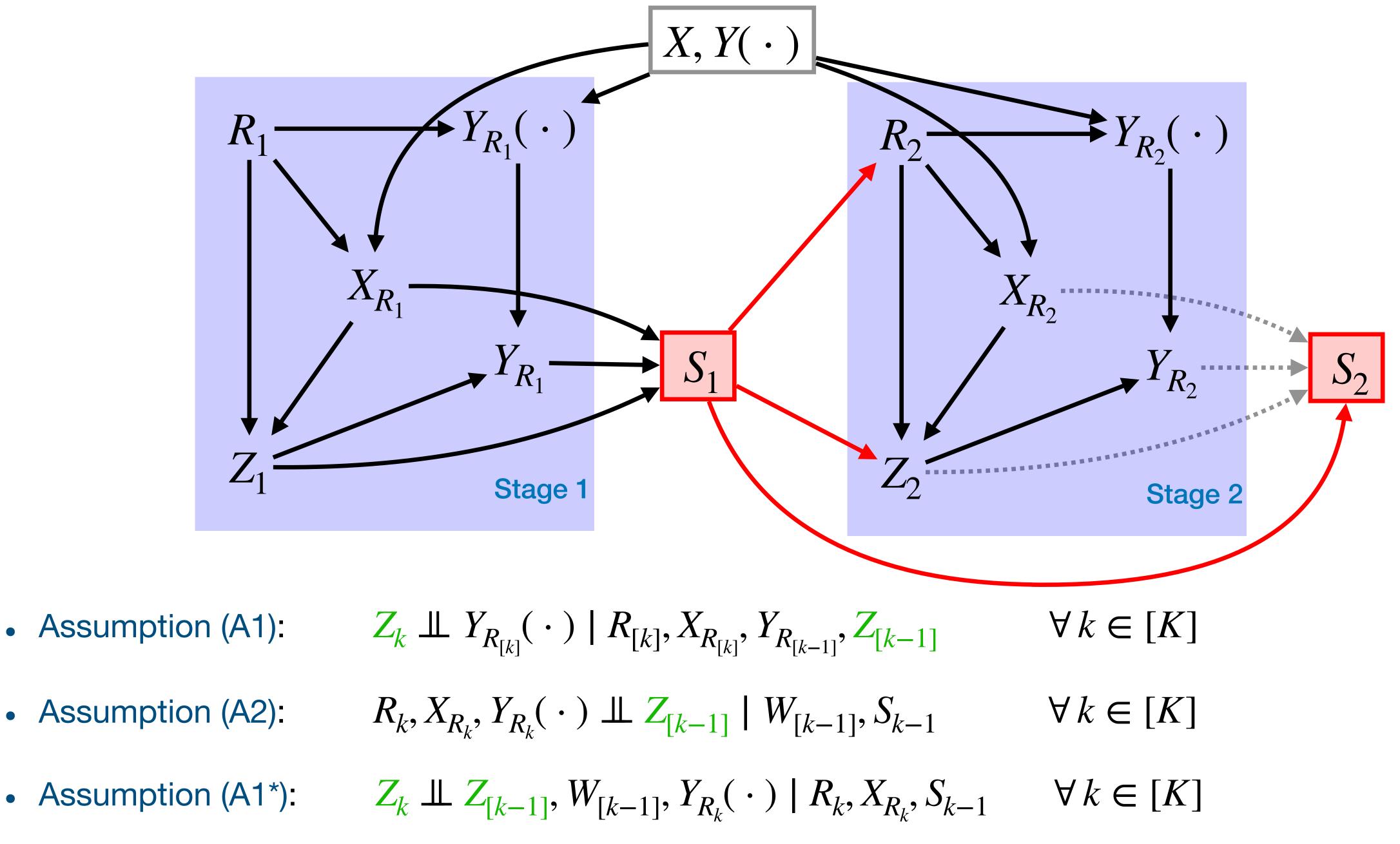
 $\cdot Y_{R_2}(\cdot)$ 

 $R_2$ 

Stage 2

\*\*\*\*\*\*\*\*\*\*\*\*\*\*





- Assumption (A1):
- Assumption (A2):

## **Randomization Inference**

- Strength: no modelling assumptions, no i.i.d. data
- Distribution of  $Z = (Z_1, \ldots, Z_K)$  is known
- Null hypothesis:  $Y_i(1) Y_i(0) = 0$  for all/subset of units
- Condition on W and compare observed value of statistic T(Z, W) against values  $T(Z^*, W)$  under alternative treatment assignments  $Z^*$ .
- $\mathbb{P}(T(Z^*, W) \leq T(Z, W) \mid W, Z)$ , wh
- Is there a problem when the experiment is adaptive?

here 
$$Z^* \stackrel{D}{=} Z$$
 and  $Z^* \perp \!\!\!\perp Z \mid W$ 

Fisher (1935), Pitman (1937), Zhang & Zhao (2023)



## **Selective Randomization Inference**

- Using data twice (double dipping)
- Comparing to  $Z^*$  that choose different stage-II design or null hypothesis than Z
- Result: Type-I error inflation
- Solutions:
  - Data splitting (Cox, 1975):

  - Selective randomization inference:

$$P_{sel} = \mathbb{P}(T(Z^*, W) \le T(Z, W) \mid W, Z, S(Z^*) = S(Z))$$

#### $\mathbb{P}(T(Z^*, W) \le T(Z, W) \mid W, Z, Z_1^* = Z_1)$ , where K = 2

Selective inference (Lee et al., 2016; Fithian et al., 2017): regression models etc.



## Computability

- computed.
- Formula for the selective randomization distribution under (A1\*):

• 
$$\mathbb{P}(Z = z \mid W = w, S(Z) = s) = \overline{\Sigma}$$
  
•  $q(z \mid w, s) = \mathbf{1}\{S(z) = s\} \cdot \prod_{k=1}^{K} \mathbb{P}(z)$   
 $\sum_{k=1}^{K} \mathbf{1}\{T(z) \in S\}$ 

Under Assumptions (A1) and (A2), the selective randomization p-value can be

 $\frac{q(z \mid w, s)}{\sum_{z'} q(z' \mid w, s)}, \text{ where }$  $(Z_k = z_k | R_k = r_k, X_{R_k} = x_{R_k}, S_{k-1} = s_{k-1})$ • Formula for p-value:  $P_{sel} = \frac{\sum_{z^*} \mathbf{1} \{ T(z^*, W) \le T(Z, W) \} q(z^* \mid W, S(Z)) }{}$ 

 $\sum_{z^*} q(z^* \mid W, S(Z))$ 

## Computation

$$P_{sel} = \mathbb{P}(T(Z^*, W) \le T$$

Monte Carlo approximation: Generate M feasible samples  $(z_i^*)_{i=1}^M$ , i.e.  $S(z_i^*) = S(Z)$ , and compute

$$\hat{P}_M := \frac{1 + \sum_{j=1}^M \mathbf{1}}{-1}$$

Rejection sampling, Markov Chain Monte Carlo (MCMC)

 $T(Z, W) \mid W, Z, S(Z^*) = S(Z))$ 

## $\{T(z_j^*, W) \le T(Z, W)\}$

1 + M

#### Inference

$$P_{sel} = \mathbb{P}(T(Z^*, W) \le T$$

- Confidence intervals:
  - test  $Y_i(1) Y_i(0) = \tau$  for different  $\tau$
  - $(1 \alpha)$  confidence interval:  $C_{1-\alpha} = \{\tau : P_{sel}(\tau) \ge \alpha\}$
- Estimation:  $\tau$  such that  $P_{sel}(\tau) = 0.5$
- Data carving: non-adaptive hold-out units

#### $\Gamma(Z, W) \mid W, Z, S(Z^*) = S(Z))$

## Simulation Study

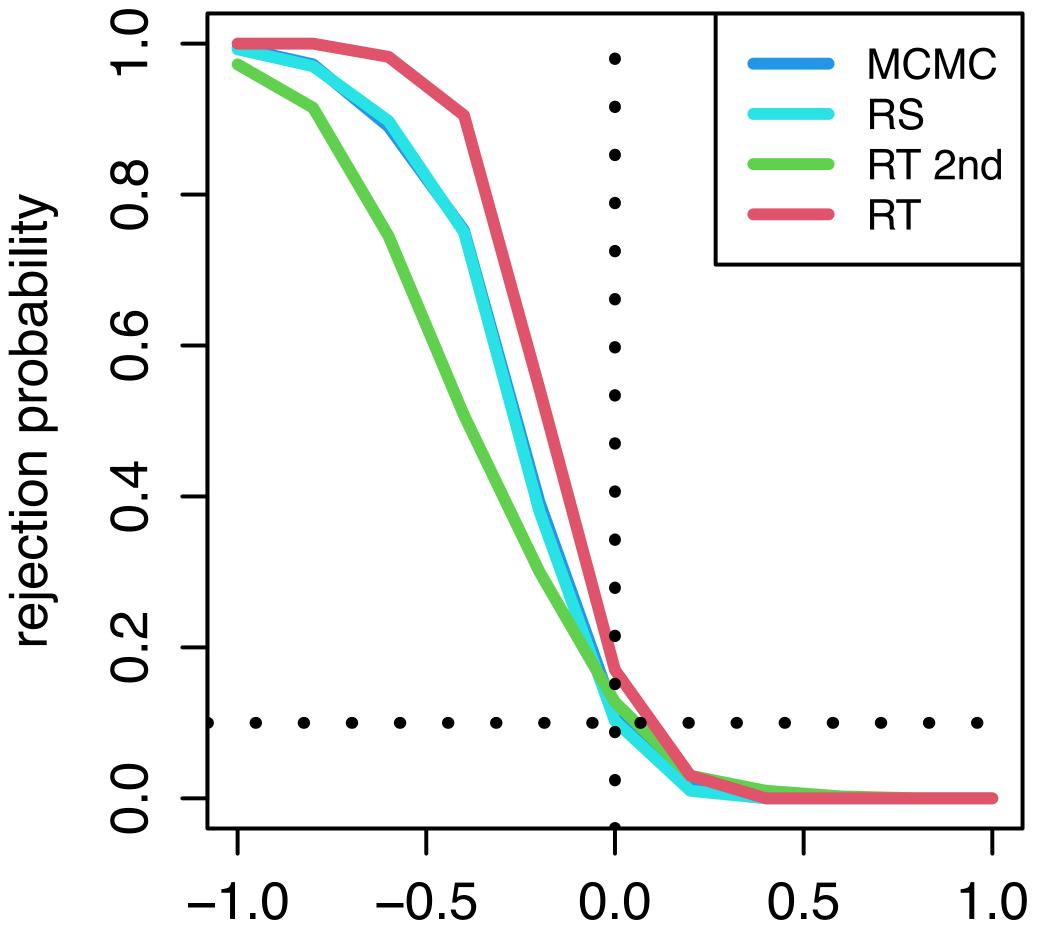
- 2 stages, 2 treatments  $Z_i \in \{0,1\}$ , 2 groups  $X_i \in \{\text{low}, \text{high}\}$
- Potential outcomes:  $Y_i(0) = Y_i(1) \sim N(0,1)$  i.i.d.
- First stage: 100 patients, Second stage: 40 patients
- $\Delta = \text{standardized difference in SATEs between groups}$
- Selection variable:

$$S = \begin{cases} \text{only low,} & \Delta < \Phi^{-1}(0.2), \\ \text{only high,} & \Delta > \Phi^{-1}(0.8), \\ \text{both,} & \text{otherwise,} \end{cases}$$

recruit 40 from group  $X_i = low$ recruit 40 from group  $X_i = high$ recruit 20 from each group

## **Power Analysis**

#### unconditional



#### • RT: n type error control MCMC RS

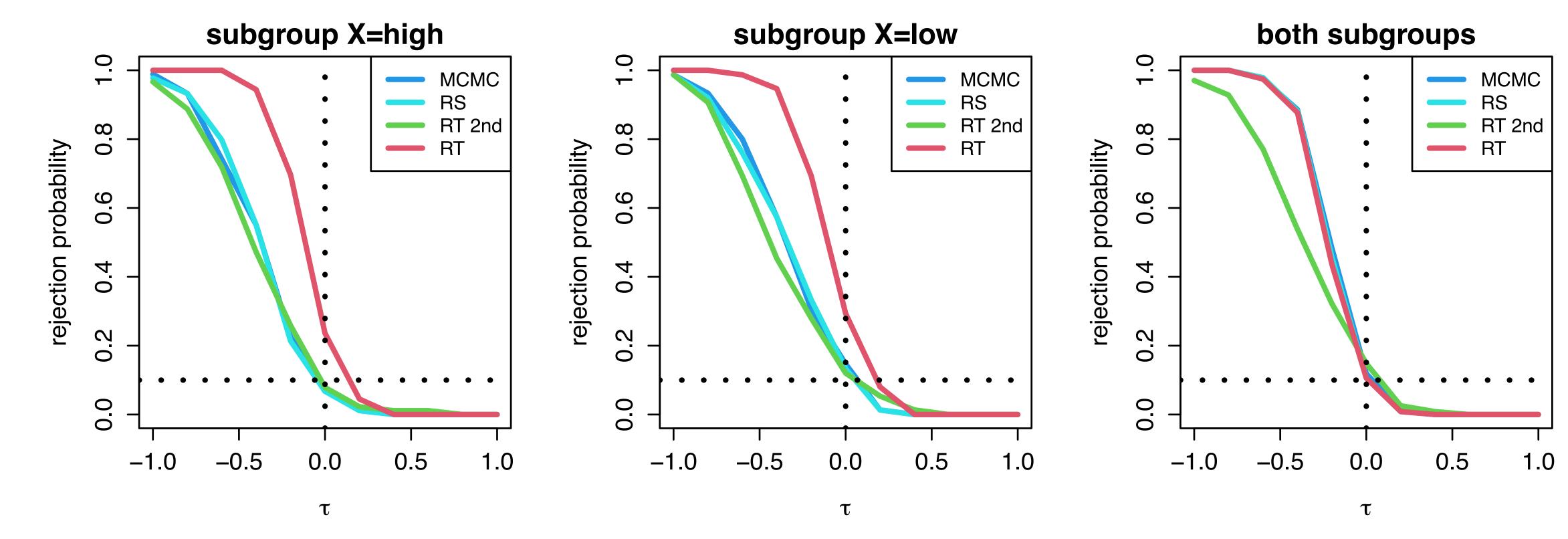
- RT 2nd: v vid but has low porrer
- Selective R<sup>1</sup> valid and more powerful.
- Rejection sampling and MCMC lead to very similar approximations.



#### **Power Analysis**

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nd



- Type-I error control in every subgroup
- Gain in power when there is a lot of "randomness left"

## Conclusion

- Experiments with adaptive treatments, recruitment and null hypothesis
- Visualization via DAGs
- Computability under general assumptions
- Approximation via rejection sampling or MCMC

#### Key idea: Conditioning randomization p-value on the selection information

# Thanks for your attention!



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#### References

Cox, D.R. (1975) 'A note on data-splitting for the evaluation of significance levels', Biometrika, 62(2), pp. 441– 444.

Fisher, R. A. (1935). 'The design of experiments', Edinburgh: Oliver & Boyd.

Fithian, W., Sun, D. and Taylor, J. (2017) 'Optimal Inference After Model Selection', arXiv:1410.2597

Lee, J.D., Sun, D.L., Sun, Y., Taylor J. (2016) 'Exact post-selection inference, with application to the lasso', The Annals of Statistics, 44(3).

Marston, N.A. et al. (2020) 'Predicting Benefit From Evolocumab Therapy in Patients With Atherosclerotic Disease Using a Genetic Risk Score', Circulation, 141(8), pp. 616–623.

Pitman, E.J.G. (1937) 'Significance Tests Which May be Applied to Samples From any Populations', Supplement to the Journal of the Royal Statistical Society, 4(1), pp. 119–130.

Zhang, Y. and Zhao, Q. (2023) 'What is a Randomization Test?', Journal of the American Statistical Association, O(0), pp. 1–15.



## **Hold-out Units**

