Selective Randomization Inference for Adaptive Experiments

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Randomized controlled trial





- Randomized controlled trial
- Finite population perspective





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- Inference without modelling- or i.i.d. data- assumptions





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Realized outcomes: $Y_i = Y_i(Z_i)$

Fisher (1935), Pitman (1937), Zhang & Zhao (2023), Imbens & Rubin (2015)

 \rightarrow



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- Potential outcomes: $Y_i(0), Y_i(1)$
- Distribution of Z is known and $Z \perp Y(\cdot) \mid X$



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İ	Υ	Y(0)	Y(1)	Ζ
1	5		5	1
2	7	7		0
3	-3	-3		0
4	0		0	1





- Null hypothesis: $Y_i(1) Y_i(0) = 0$ for all *i*
- Test statistic: $T(Z, Y(\cdot))$, e.g. average outcome of treated minus control

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- P-value:

 $\mathbb{P}(T(Z^*, Y(\cdot)) \leq T(Z, Y(\cdot)) \mid Y(\cdot), Z),$ where $Z^* \stackrel{D}{=} Z$ and $Z^* \perp Z \mid Y(\cdot)$

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Example









Example

















Graphical Model

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Pearl (2009)



Graphical Model

$$X, Y(\cdot)$$

- Covariates: *X*
- Potential outcomes: $Y(\cdot)$





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 S_2

- Observed outcomes: Y = Y(Z)
- Selective choice: S_k

 X_{R_2}

 $\bullet Y_{R_2}(\cdot)$

 R_2

Stage 2







- Covariates: X
- Potential outcomes: $Y(\cdot)$
- Recruitment: $R_k \subseteq [n]$
- Treatments: Z_k

 S_2

- Observed outcomes: Y = Y(Z)
- Selective choice: S_k
- Short-hand: $W = (R, X_R, Y_R(\cdot))$

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 $Y_{R_2}(\cdot)$

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Stage 2









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- Assumption (A2): Conditional independence of present and post potential outcomes $Z_{k} \perp Y_{R_{[k]}}(\cdot) \mid R_{[k]}, X_{R_{[k]}}, Y_{R_{[k-1]}}, Z_{[k-1]} \quad \forall k \in [K]$



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- Assumption (A3): S_k captures all the information that is passed on $R_k, X_{R_k}, Y_{R_k}(\cdot) \perp Z_{[k-1]} \mid W_{[k-1]}, S_{k-1} \quad \forall k \in [K]$

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 - Data splitting (Cox, 1975):

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Selective inference (Lee et al., 2016; Fithian et al., 2017): regression models etc.



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 - Selective inference (Lee et al., 2016; Fithian et al., 2017): regression models etc.
 - Selective randomization inference:

$$P_{sel} = \mathbb{P}(T(Z^*, W) \le T(Z, W) \mid W, Z, S(Z^*) = S(Z))$$



Type-I Error

$$P_{sel} = \mathbb{P}(T(Z^*, W) \le T$$

 $T(Z, W) \mid W, Z, S(Z^*) = S(Z))$

Type-I Error

$$P_{sel} = \mathbb{P}(T(Z^*, W) \le T$$

• **Proposition:** P_{sel} controls the selective type-I error, i.e.

 $\mathbb{P}(P_{sel} \le \alpha \mid W, S(Z) = s) \le \alpha$

 $V(Z, W) \mid W, Z, S(Z^*) = S(Z))$

 $\forall s$

Type-I Error

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- **Proposition:** P_{sel} controls the selective type-I error, i.e. $\mathbb{P}(P_{sel} \leq \alpha \mid W, \mathcal{L})$ • Formula for p-value: $P_{sel} = \frac{\sum_{z^*} \mathbf{1} \{ T(z^*,$

•
$$q(z \mid w) = \prod_{k=1}^{K} \mathbb{P}(Z_k = z_k \mid R_k = r_k, X_{R_k})$$

$$S(Z) = s) \le \alpha \qquad \forall s$$

$$(W) \leq T(Z, W) \} \mathbf{1} \{ S(z^*) = S(Z) \} q(z^* | W)$$

$$\sum_{z^*} \mathbf{1} \{ S(z^*) = S(Z) \} q(z^* \mid W)$$

 $= x_{R_k}, S_{k-1} = s_{k-1}$) under slightly stronger assumptions

Computation

$$P_{sel} = \mathbb{P}(T(Z^*, W) \le T$$

$T(Z, W) \mid W, Z, S(Z^*) = S(Z))$

Computation

$$P_{sel} = \mathbb{P}(T(Z^*, W) \le T$$

• Monte Carlo approximation: Generate M feasible samples $(z_j^*)_{j=1}^M$, i.e. $S(z_j^*) = S(Z)$, and compute

$$\hat{P}_M := \frac{1 + \sum_{j=1}^M \mathbf{1}}{2}$$

 $\Gamma(Z, W) \mid W, Z, S(Z^*) = S(Z))$

 $\{T(z_j^*, W) \le T(Z, W)\}$

1 + M

Computation

$$P_{sel} = \mathbb{P}(T(Z^*, W) \le T$$

Monte Carlo approximation: Generate M feasible samples $(z_i^*)_{i=1}^M$, i.e. $S(z_i^*) = S(Z)$, and compute

$$\hat{P}_M := \frac{1 + \sum_{j=1}^M \mathbf{1}}{-1}$$

Rejection sampling, Markov Chain Monte Carlo (MCMC)

 $T(Z, W) \mid W, Z, S(Z^*) = S(Z))$

$\{T(z_j^*, W) \le T(Z, W)\}$

1 + M

Inference

$$P_{sel} = \mathbb{P}(T(Z^*, W) \le T$$

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Inference

$$P_{sel} = \mathbb{P}(T(Z^*, W) \le T$$

• Confidence intervals:

$T(Z, W) \mid W, Z, S(Z^*) = S(Z))$

Inference

$$P_{sel} = \mathbb{P}(T(Z^*, W) \le T$$

- Confidence intervals:
 - test $Y_i(1) Y_i(0) = \tau$ for different τ
 - (1α) confidence interval: $C_{1-\alpha} = \{\tau : P_{sel}(\tau) \ge \alpha\}$
- Estimation: τ such that $P_{sel}(\tau) = 0.5$

$T(Z, W) \mid W, Z, S(Z^*) = S(Z))$

- 2 stages, 2 treatments $Z_i \in \{0,1\}$, 2 groups $X_i \in \{low, high\}$
- Potential outcomes: $Y_i(0) = Y_i(1) \sim N(0,1)$ i.i.d.
- First stage: 100 patients, Second stage: 40 patients

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- First stage: 100 patients, Second stage: 40 patients
- $\Delta = \text{standardized difference in SATEs between groups}$
- Selection variable:

$$S = \begin{cases} \text{only low,} & \Delta < \Phi^{-1}(0.2), \\ \text{only high,} & \Delta > \Phi^{-1}(0.8), \\ \text{both,} & \text{otherwise,} \end{cases}$$

recruit 40 from group $X_i = low$ recruit 40 from group $X_i = high$ recruit 20 from each group

unconditional







unconditional







unconditional





unconditional



• RT: n type error control MCMC RS

- RT 2nd: v vid but has low porrer
- Selective R¹ valid and more powerful.
- Rejection sampling and MCMC lead to very similar approximations.



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• Type-I error control in every subgroup

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- Type-I error control in every subgroup
- Gain in power when there is a lot of "randomness left"

Conclusion

- Robust inference in randomized controlled trials
- Experiments with adaptive treatments, recruitment and null hypothesis
- Visualization via graphs
- Approximation via rejection sampling or MCMC

Key idea: Conditioning randomization p-value on the selection information

Thanks for your attention!



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Hold-out Units

