

# Bayesian Optimal Experimental Design of Clinical Studies

Tobias Freidling

DPMMS, University of Cambridge

March 22, 2020

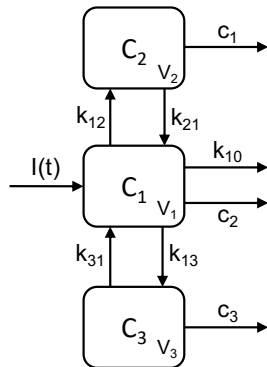
supervised by Thomas Moxon, Unilever

# Physiologically based pharmacokinetic (PBPK) model

$$V_1 \frac{dC_1}{dt} = - [c_1 + V_1(k_{12} + k_{13} + k_{10})] C_1 + k_{21} V_2 C_2 + k_{31} V_3 C_3 + I(t)$$

$$V_2 \frac{dC_2}{dt} = - [c_2 + k_{21} V_2] C_2 + k_{12} V_1 C_1$$

$$V_3 \frac{dC_3}{dt} = - [c_3 + k_{31} V_3] C_3 + k_{13} V_1 C_1$$



**Figure:** Schematic of the 3-compartment model for Remifentanyl

# PBPK model

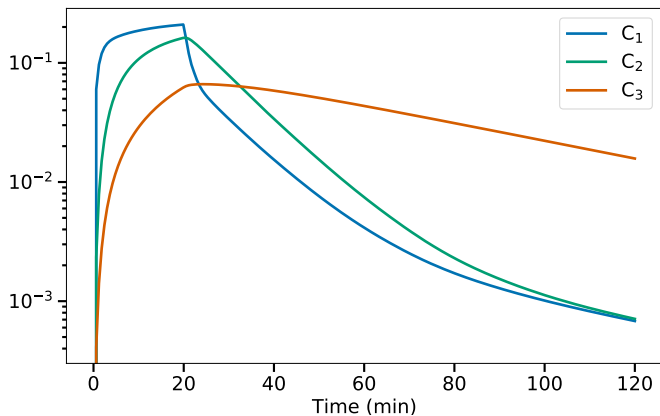


Figure: Concentrations in the 3 compartments

# Bayesian 3-Compartment Model

Let  $d = (t_1, \dots, t_p)$ ,  $\bar{\theta} = (k_{10}, k_{12}, \dots)$  and  $C(d, \bar{\theta}) \in \mathbb{R}^{3p}$  the solution of the ODE system evaluated at  $d$ .

# Bayesian 3-Compartment Model

Let  $d = (t_1, \dots, t_p)$ ,  $\bar{\theta} = (k_{10}, k_{12}, \dots)$  and  $C(d, \bar{\theta}) \in \mathbb{R}^{3p}$  the solution of the ODE system evaluated at  $d$ .

Bayesian model:

$$p(y|\theta, d) \sim \mathcal{N}(\log C(d, \bar{\theta}), \sigma^2 \text{Id}), \quad p(\sigma^2) \sim \text{Gamma}(2, 31),$$

Gaussian prior for volumes ( $V$ -parameters) and flow rates ( $k$ -parameters), and log-normal prior for clearance rates ( $c$ -parameters). We denote the prior distribution  $p(\theta)$ , where  $\theta = (\bar{\theta}, \sigma^2)$ .

Expected information gain (EIG):

$$\begin{aligned} \text{EIG}(d) &:= \mathbb{E}_{p(y|d)} \left[ D_{\text{KL}} (p(\theta|y, d) \parallel p(\theta)) \right] \\ &= \mathbb{E}_{p(y, \theta|d)} \left[ \log \frac{p(\theta|y, d)}{p(\theta)} \right] = \mathbb{E}_{p(y, \theta|d)} \left[ \log \frac{p(y|\theta, d)}{p(y|d)} \right]. \end{aligned}$$

Expected information gain (EIG):

$$\begin{aligned}\text{EIG}(d) &:= \mathbb{E}_{p(y|d)} \left[ D_{\text{KL}} (p(\theta|y, d) \parallel p(\theta)) \right] \\ &= \mathbb{E}_{p(y, \theta|d)} \left[ \log \frac{p(\theta|y, d)}{p(\theta)} \right] = \mathbb{E}_{p(y, \theta|d)} \left[ \log \frac{p(y|\theta, d)}{p(y|d)} \right].\end{aligned}$$

Nested Monte Carlo (NMC) estimator:

$$\hat{\mu}_{\text{NMC}}(d) = \frac{1}{N} \sum_{n=1}^N \log \frac{p(y_n | \theta_{n,0}, d)}{\frac{1}{M} \sum_{m=1}^M p(y_n | \theta_{n,m}, d)},$$

with iid samples  $\theta_{n,m} \sim p(\theta)$  and  $y_n \sim p(y|\theta_{n,0}, d)$ .

NMC converges to EIG from above but is computationally expensive.

# Variational Marginal

We approximate  $p(y|d)$  by  $q(y|d, \phi)$  and compute the estimate

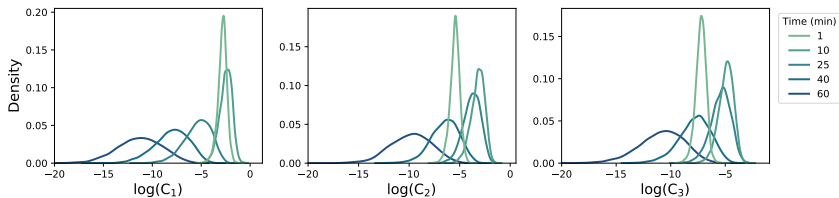
$$\hat{\mu}_{\text{marg}}(d) = \frac{1}{N} \sum_{n=1}^N \log \frac{p(y_n|\theta_n, d)}{q(y_n|d, \phi)},$$

where  $y_n, \theta_n \sim p(y, \theta|d)$  is an iid sample.

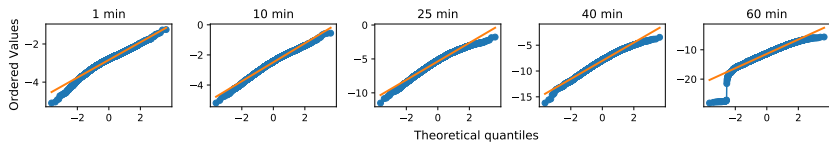
The variational marginal estimate is asymptotically an upper bound to EIG that is sharp if and only if  $q(y|d, \phi) = p(y|d)$ .



# Marginal distribution



(a) Distribution of the log-concentrations in the three compartments



(b) P-P plot of a Gaussian distribution against  $\log(C_1)$

# Gaussian and KDE approximation

**Gaussian:**  $q(y|d, \phi) \sim \mathcal{N}(\mu, \Sigma), \phi = (\mu, \Sigma)$

We simulate data points from the model, estimate  $\hat{\mu}$  and  $\hat{\Sigma}$  and use  $q(y|d, \phi) \sim \mathcal{N}(\hat{\mu}, \hat{\Sigma})$ .

# Gaussian and KDE approximation

**Gaussian:**  $q(y|d, \phi) \sim \mathcal{N}(\mu, \Sigma), \phi = (\mu, \Sigma)$

We simulate data points from the model, estimate  $\hat{\mu}$  and  $\hat{\Sigma}$  and use  $q(y|d, \phi) \sim \mathcal{N}(\hat{\mu}, \hat{\Sigma})$ .

**Kernel density estimation (KDE):**

$$q(y|d, \phi) := \frac{1}{n \det(H)} \sum_{i=1}^n K(H^{-1}(y - Y_i)),$$

where  $Y_1, \dots, Y_n$  are samples from the model.  $K$  is a nonnegative function that integrates to 1 and  $H$  is the bandwidth matrix.

# Gaussian and KDE approximation

**Gaussian:**  $q(y|d, \phi) \sim \mathcal{N}(\mu, \Sigma), \phi = (\mu, \Sigma)$

We simulate data points from the model, estimate  $\hat{\mu}$  and  $\hat{\Sigma}$  and use  $q(y|d, \phi) \sim \mathcal{N}(\hat{\mu}, \hat{\Sigma})$ .

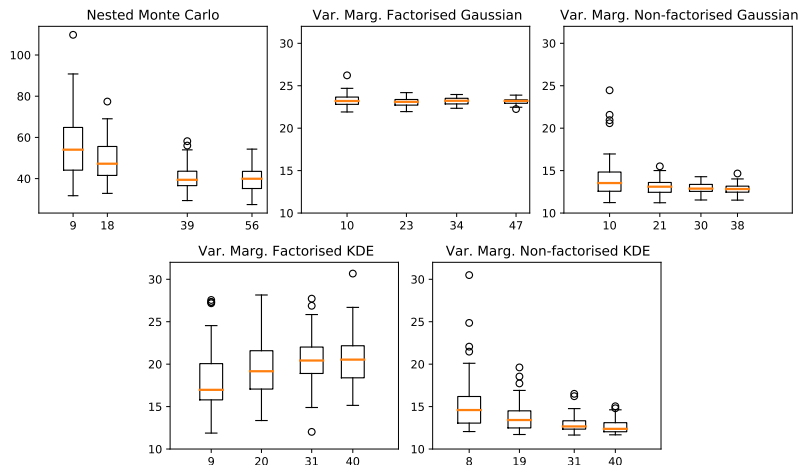
**Kernel density estimation (KDE):**

$$q(y|d, \phi) := \frac{1}{n \det(H)} \sum_{i=1}^n K(H^{-1}(y - Y_i)),$$

where  $Y_1, \dots, Y_n$  are samples from the model.  $K$  is a nonnegative function that integrates to 1 and  $H$  is the bandwidth matrix.

We consider factorised (diagonal  $\Sigma$  and  $H$ ) and non-factorised approximations.

# Empirical Evaluation of Estimators



**Figure:** Box plot of different EIG-estimators for the design  $d = (1, 10, 25, 40, 60)$ ; average runtimes [s] on x-axis

# Optimisation

- ▶ Constraints on design  $d$ : at least 5 minutes between two measurements
- ▶ No gradient information  $\nabla_d \text{EIG} \rightarrow$  Simulated Annealing or Differential Evolution
- ▶ Unstable estimates of EIG 😞
- ▶ Grid search or list of test designs

EIG	Design	EIG	Design
13.69	(30, 55, 60)	11.93	(5, 45, 60)
13.02	(35, 55, 60)	11.87	(20, 50, 60)
12.75	(35, 45, 60)	11.86	(20, 55, 60)
12.04	(30, 50, 60)	11.85	(40, 55, 60)
11.93	(50, 55, 60)	11.84	(30, 50, 55)

**Table:** The 10 best designs for three measurements according to grid search; variational marginal estimator with non-factorised Gaussian approximation

- ▶ Hyper-parameter tuning and sensitivity w.r.t. prior distributions
- ▶ Multiple participants with potentially different designs
- ▶ New EIG-estimators
- ▶ EIG estimation for a subset of parameters
- ▶ Optimisation algorithm for instable function evaluations

Thank you!